

Uniform Estimate for Distributions of the Sum of i.i.d. Random Variables with Fat Tail

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Abstract. The research on asymptotic behavior of distributions of the sum of i.i.d random variables has a long history and a lot of facts are known. The authors consider the case where the distribution of a random variable has the second moment but has a fat tail, and they show a new limit theorem for large deviations.

1. Introduction

Let (Ω, \mathcal{F}, P) be a probability space and $X_n, n = 1, 2, \dots$, be independent identically distributed random variables with the same probability law μ .

In the present paper we assume that

(A-1) $E[X_1^2] = 1$ and $E[X_1] = 0$.

Let $F : \mathbf{R} \rightarrow [0, 1]$ and $\bar{F} : \mathbf{R} \rightarrow [0, 1]$ be given by

$$F(x) = \mu((-\infty, x]) = P(X_1 \leq x) \text{ and} \\ \bar{F}(x) = \mu((x, \infty)) = P(X_1 > x), \quad x \in \mathbf{R}.$$

We also assume the following.

(A-2) $\bar{F}(x)$ is a regularly varying function of index $-\alpha$ for some $\alpha > 2$, as $x \rightarrow \infty$, i.e., if we let

$$L(x) = x^\alpha \bar{F}(x), \quad x \geq 1,$$

then $L(x) > 0$ for any $x \geq 1$, and for any $a > 0$

$$\frac{L(ax)}{L(x)} \rightarrow 1, \quad x \rightarrow \infty.$$

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$$(A-3) \quad |x|^{\alpha+2}F(x) \rightarrow 0, \quad x \rightarrow -\infty.$$

Recently people in finance are interested in computing the quantile of the distribution of $\sum_{k=1}^n X_k$ for the purpose of measuring market risk.

There are many works on this topic. In particular, there are many results on large deviation results (e.g. Borovkov-Borovkov [1], also see books, Borovkov-Borovkov [2] and Petrov [7]). However, there are not so many results on uniform estimates. Nagaev [5] and [6] proved the following theorem (also see Linnik [4]), and this is the best result so far to our best knowledge.

THEOREM 1 (Nagaev). *Assume (A-1)-(A-3). Then we have*

$$\sup_{s \in [1, \infty)} \left| \frac{P(\sum_{k=1}^n X_k > sn^{1/2})}{\Phi_0(s) + n\bar{F}(sn^{1/2})} - 1 \right| \rightarrow 0, \quad n \rightarrow \infty.$$

Here $\Phi_0 : \mathbf{R} \rightarrow \mathbf{R}$ is given by

$$\Phi_0(x) = \frac{1}{\sqrt{2\pi}} \int_x^\infty \exp\left(-\frac{y^2}{2}\right) dy, \quad x \in \mathbf{R}.$$

In this paper, we show two theorems. Combining them, we can improve Nagaev's result a little bit.

Let us explain our results. We assume the following assumption furthermore.

(A-4) The probability law μ is absolutely continuous and has a density function $\rho : \mathbf{R} \rightarrow [0, \infty)$ which is right continuous and has a finite total variation.

To state our theorem (Theorem 2), we need some preparations.

Let K be an integer such that $K - 1 < \alpha \leq K$. Then $K \geq 3$. From the assumptions (A-2) and (A-3), we see that the probability law μ has $(K - 1)$ -th moment. So let η_k , $k = 1, \dots, K - 1$, be given by

$$\eta_k = \int_{\mathbf{R}} x^k \mu(dx).$$

Then we see that $\eta_1 = 0$ and $\eta_2 = 1$. Also, let us define $\Phi_k : \mathbf{R} \rightarrow \mathbf{R}$, $k = 1, 2, \dots$, by

$$\Phi_1(x) = \frac{1}{\sqrt{2\pi}} \exp\left(-\frac{x^2}{2}\right) = -\frac{d}{dx} \Phi_0(x),$$

and

$$\Phi_k(x) = (-1)^{k-1} \frac{d^{k-1}}{dx^{k-1}} \Phi_1(x), \quad k = 2, 3, \dots$$

THEOREM 2. *Assume (A-1)-(A-4). Then there are $\delta > 0$ and $C > 0$ such that*

$$\sup_{s \in [1, \infty)} |P(\sum_{k=1}^n X_k > sn^{1/2}) - G(n, s)| \leq Cn^{-(\alpha-2)/2-\delta}, \quad n = 3, 4, \dots$$

Here

$$\begin{aligned} & G(n, s) \\ = & \Phi_0(s) + n \int_{-\infty}^s \bar{F}((s-x)n^{1/2}) \Phi_1(x) dx - \sum_{k=1}^{K-1} \frac{n^{-(k-2)/2}}{k!} \Phi_k(s) \int_0^\infty x^k \mu(dx) \\ & + \frac{n^{-(K-2)/2}}{K!} \Phi_K(s) \int_{-\infty}^0 x^K \mu(dx) + \sum_{k=3}^{K-1} n^{-k/6} q_k(n^{-1/3}, \eta_2, \dots, \eta_k) \Phi_k(s) \\ & + \sum_{k=K}^{3(K-1)} n^{-k/6} q_k(n^{-1/3}, \eta_2, \dots, \eta_{K-1}, 0, \dots, 0) \Phi_k(s), \end{aligned}$$

and q_k 's are polynomials defined in the next section.

For the next theorem we assume the following also.

(A-5) There is an $x_0 > 0$ such that \bar{F} is twice continuously differentiable on (x_0, ∞) and that

$$x^2 \frac{d^2}{dx^2} \log \bar{F}(x) \rightarrow \alpha, \quad x \rightarrow \infty.$$

Then we have the following.

THEOREM 3. *Assume the assumptions (A-1)- (A-5) and let $\beta : \mathbf{N} \rightarrow (0, \infty)$ be such that*

$$\frac{\beta(n)}{(\log n)^{1/2}} \rightarrow \infty, \quad n \rightarrow \infty.$$

Then we have

$$\sup_{s \geq n^{1/2}\beta(n)} \frac{s^2}{n} \left| \frac{P(\sum_{k=1}^n X_k > s)}{n\bar{F}(s)} - \left(1 + \frac{\alpha(\alpha+1)n}{2s^2}\right) \right| \rightarrow 0, \quad n \rightarrow \infty.$$

Let

$$H(n, s) = \Phi_0(s) + n \int_{-\infty}^s \bar{F}((s-x)n^{1/2})\Phi_1(x)dx - \sum_{k=1}^2 \frac{n^{-(k-2)/2}}{k!} \Phi_k(s) \int_0^\infty x^k \mu(dx)$$

for $s \geq 1$, and $n \geq 1$.

Then we also show the following.

THEOREM 4. *Assume (A-1)-(A-5). Then there exist a $C > 0$, $\delta > 0$ and $n_0 \geq 1$ such that*

$$\sup_{s \in [1, \infty)} \left| \frac{P(\sum_{k=1}^n X_k > sn^{1/2})}{H(n, s)} - 1 \right| \leq Cn^{-\delta}, \quad n \geq n_0.$$

Note that by Theorem 3, we see that

$$2(\log n)^2 \left(\frac{P(\sum_{k=1}^n X_k > (\log n)n^{1/2})}{\Phi_0(\log n) + n\bar{F}((\log n)n^{1/2})} - 1 \right) \rightarrow \alpha(\alpha+1) \quad n \rightarrow \infty.$$

Therefore we see that $H(n, s)$ is a better approximation for $P(\sum_{k=1}^n X_k > sn^{1/2})$ than $\Phi_0(s) + n\bar{F}(sn^{1/2})$.

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2. Algebraic Preparation

In this section, we think of formal power series in z . First, we think of the following formal power series in z .

$$(1) \quad \log\left(1 + \sum_{k=2}^{\infty} \frac{a_k}{k!} z^k\right) = \sum_{\ell=1}^{\infty} \frac{(-1)^{\ell-1}}{\ell} \left(\sum_{k=2}^{\infty} \frac{a_k}{k!} z^k\right)^\ell = \sum_{\ell=2}^{\infty} c_\ell(a_2, \dots, a_\ell) \frac{z^\ell}{\ell!}$$

Then we see that $c_\ell(a_2, \dots, a_\ell)$, $\ell \geq 2$, are polynomials in a_2, \dots, a_ℓ , and

$$c_\ell(t^2 a_2, \dots, t^\ell a_\ell) = t^\ell c_\ell(a_2, \dots, a_\ell)$$

for any $t, a_1, \dots, a_\ell \in \mathbf{R}$. Moreover, we see that

$$c_2(a_2) = a_2 \quad \text{and} \quad c_\ell(a_2, \dots, a_{\ell-1}, a_\ell) = c_\ell(a_2, \dots, a_{\ell-1}, 0) + a_\ell, \quad \ell \geq 2.$$

We also think of the following formal power series in z .

$$\begin{aligned} & \exp\left(y^{-3} \sum_{\ell=3}^{\infty} c_\ell(a_2, \dots, a_\ell) \frac{(yz)^\ell}{\ell!}\right) \\ (2) \quad & = 1 + \sum_{k=1}^{\infty} \frac{1}{k!} \left(\sum_{\ell=3}^{\infty} c_\ell(a_2, \dots, a_\ell) \frac{y^{\ell-3} z^\ell}{\ell!} \right)^k = 1 + \sum_{k=3}^{\infty} q_k(y, a_2, \dots, a_k) z^k. \end{aligned}$$

Then we see that $q_k(y, a_2, \dots, a_k)$, $k \geq 3$, are polynomials in y, a_2, \dots, a_ℓ . Note that

$$q_k(y, t^2 a_2 \dots, t^k a_k) = t^k q_k(y, a_2, \dots, a_k)$$

and that

$$q_k(y, a_2, \dots, a_k) = q_k(y, a_2, \dots, a_{k-1}, 0) + \frac{y^{k-3}}{k!} a_k, \quad k \geq 3.$$

Also we have

$$\begin{aligned} & \exp\left(y^{-6} \sum_{\ell=3}^{\infty} c_\ell(a_2, \dots, a_\ell) \frac{(y^3 z)^\ell}{\ell!}\right) \\ & = \exp\left((y^2)^{-3} \sum_{\ell=3}^{\infty} c_\ell(y^2 a_2, \dots, y^\ell a_\ell) \frac{(y^2 z)^\ell}{\ell!}\right) \\ (3) \quad & = 1 + \sum_{k=3}^{\infty} q_k(y^2, y^2 a_2, \dots, y^k a_k) z^k = 1 + \sum_{k=3}^{\infty} y^k q_k(y^2, a_2, \dots, a_k) z^k \end{aligned}$$

as a formal power series in z .

3. Property of the Function L

PROPOSITION 5. *We have*

$$\sup_{1/2 \leq a \leq 2} \frac{L(ax)}{L(x)} \rightarrow 1, \quad x \rightarrow \infty,$$

and

$$\inf_{1/2 \leq a \leq 2} \frac{L(ax)}{L(x)} \rightarrow 1, \quad x \rightarrow \infty.$$

PROOF. Since the proof is similar, we prove the first equation only. If not, there are $\varepsilon > 0$, $\{a_n\}_{n=1}^{\infty}$ and $\{x_n\}_{n=1}^{\infty}$ such that $1/2 \leq a_n \leq 2$, $x_n \geq 1$, $n = 1, 2, \dots$, $x_n \rightarrow \infty$, $n \rightarrow \infty$, and that

$$\frac{L(a_n x_n)}{L(x_n)} > 1 + \varepsilon, \quad n = 1, 2, \dots$$

Then taking a subsequence if necessary, we may assume that there is an $a \in [1/2, 2]$ such that $a_n \rightarrow a$, $n \rightarrow \infty$. Then we see that for any $m \geq 3$ there is a $n(m) \geq 1$ such that

$$\begin{aligned} \left(a - \frac{1}{m}\right)^{-\alpha} L\left(\left(a - \frac{1}{m}\right)x_n\right) &= \bar{F}\left(\left(a - \frac{1}{m}\right)x_n\right) \geq \bar{F}(a_n x_n) \\ &= a_n^{-\alpha} L(a_n x_n), \quad n \geq n(m). \end{aligned}$$

So we have

$$\left(1 - \frac{1}{ma}\right)^{-\alpha} \geq \lim_{n \rightarrow \infty} \frac{L(a_n x_n)}{L\left(\left(a - \frac{1}{m}\right)x_n\right)} \geq 1 + \varepsilon, \quad m \geq 3.$$

Since m is arbitrary, this implies a contradiction. \square

PROPOSITION 6. *For any $\varepsilon \in (0, 1)$, there is an $M \geq 1$ such that*

$$M^{-1}y^{-\varepsilon} \leq \frac{L(yx)}{L(x)} \leq My^{\varepsilon} \quad x, y \geq 1.$$

PROOF. For any $\varepsilon \in (0, 1)$ there is an $m \geq 1$ such that

$$\left|\frac{L(ex)}{L(x)} - 1\right| \leq \varepsilon \quad x \geq e^m.$$

Let

$$C = \sup_{x \in [1, e^m]} \left(\frac{L(ex)}{L(x)} + \frac{L(x)}{L(ex)} \right) < \infty.$$

Then we have

$$C^{-m}(1 - \varepsilon)^n \leq \frac{L(e^n x)}{L(x)} \leq C^m(1 + \varepsilon)^n, \quad x \geq 1, \quad n \geq 0.$$

For any $y \geq 1$, there is an $n \geq 1$ such that $e^{n-1} \leq y \leq e^n$. Then we have

$$\bar{F}(e^{n-1}x) \geq \bar{F}(yx) \geq \bar{F}(e^n x).$$

So we have for any $x, y \geq 1$

$$\begin{aligned} (e^{-1}yx)^{-\alpha}L(e^{n-1}x) &\geq (e^{n-1}x)^{-\alpha}L(e^{n-1}x) \geq (yx)^{-\alpha}L(yx) \\ &\geq (e^n x)^{-\alpha}L(e^n x) \geq (eyx)^{-\alpha}L(e^n x), \end{aligned}$$

which implies

$$C^{-m}e^{-\alpha}(1 - \varepsilon)^n \leq \frac{L(yx)}{L(x)} \leq C^m e^\alpha (1 + \varepsilon)^{n-1}.$$

Therefore we have

$$C^{-m}e^{-\alpha}(1 - \varepsilon)y^{\log(1-\varepsilon)} \leq \frac{L(yx)}{L(x)} \leq C^m e^\alpha y^{\log(1+\varepsilon)}, \quad x \geq 1, \quad y \geq 1.$$

This implies our assertion. \square

The following is known as Karamata's theorem (c.f.[3] Appendix), but we give a proof.

PROPOSITION 7. (1) For any $\beta < -1$,

$$\frac{1}{t^{\beta+1}L(t)} \int_t^\infty x^\beta L(x) dx \rightarrow -\frac{1}{\beta+1}, \quad t \rightarrow \infty.$$

(2) For any $\beta > -1$,

$$\frac{1}{t^{\beta+1}L(t)} \int_1^t x^\beta L(x) dx \rightarrow \frac{1}{\beta+1}, \quad t \rightarrow \infty.$$

(3) Let $f : [1, \infty) \rightarrow (0, \infty)$ be given by

$$f(t) = \int_1^t x^{-1}L(x)dx \quad t \geq 1.$$

Then f is slowly varying.

PROOF. Note that for $t > 1$

$$\frac{1}{t^{\beta+1}L(t)} \int_t^\infty x^\beta L(x)dx = \int_1^\infty x^\beta \frac{L(tx)}{L(t)}dx, \text{ if } \beta < -1$$

and

$$\frac{1}{t^{\beta+1}L(t)} \int_1^t x^\beta L(x)dx = \int_{1/t}^1 x^\beta \left(\frac{L(t)}{L(tx)}\right)^{-1}dx \text{ if } \beta > -1$$

Then the assertions (1) and (2) follow from this equation and Proposition 5.

Let us prove (3). If $\lim_{t \rightarrow \infty} f(t) < \infty$, the assertion is obvious. So we assume that $\lim_{t \rightarrow \infty} f(t) = \infty$. Then for any $a > 0$ and $t_0 > 1$

$$f(at) = \int_{1/a}^t x^{-1}L(ax)dx = \int_{1/a}^{t_0} x^{-1}L(ax)dx + \int_{t_0}^t x^{-1}L(x) \frac{L(ax)}{L(x)}dx.$$

So we have

$$\inf_{x \geq t_0} \frac{L(ax)}{L(x)} \leq \underline{\lim}_{t \rightarrow \infty} \frac{f(at)}{f(t)} \leq \overline{\lim}_{t \rightarrow \infty} \frac{f(at)}{f(t)} \leq \sup_{x \geq t_0} \frac{L(ax)}{L(x)}.$$

Therefore by Proposition 6 and Lebesgue's convergence theorem, we have our assertion. \square

4. Estimate for Moments and Characteristic Functions

Remind that K is an integer such that $K - 1 < \alpha \leq K$ and

$$\eta_k = \int_{-\infty}^\infty x^k \mu(dx), \quad k = 1, 2, \dots, K - 1.$$

Then by the assumption (A4) we have $\eta_1 = 0$ and $\eta_2 = 1$. Note that

$$1 - \bar{F}(t) \geq 1 - \int_2^\infty \frac{x^2}{4} \mu(dx) \geq \frac{3}{4}$$

for any $t \geq 2$. Let

$$\eta_k(t) = \int_{(-\infty, t]} x^k \mu(dx), \quad t > 0, \quad k = 1, 2, \dots, K + 1,$$

and

$$\bar{\eta}_k(t) = \int_{(t, \infty)} x^k \mu(dx), \quad t > 0, \quad k = 0, 1, 2, \dots, K - 1.$$

Then we have

$$\eta_k(t) = \int_{(-\infty, 0)} x^k \mu(dx) + k \int_0^t x^{k-1} \bar{F}(x) dx - t^k \bar{F}(t),$$

$$t > 0, \quad k = 1, 2, \dots, K + 1,$$

and

$$\bar{\eta}_k(t) = k \int_t^\infty x^{k-1} \bar{F}(x) dx + t^k \bar{F}(t) \quad t > 0, \quad k = 0, 1, 2, \dots, K - 1.$$

Then by Propositions 6 and 7 we have the following.

PROPOSITION 8. *For any $\varepsilon > 0$, there is a $C(\varepsilon) > 0$ such that*

$$L(t) \leq C(\varepsilon)t^\varepsilon,$$

$$|\eta_K(t)| \leq C(\varepsilon)t^{-\alpha+K+\varepsilon},$$

$$|\bar{\eta}_k(t)| \leq C(\varepsilon)t^{-\alpha+k+\varepsilon}, \quad k = 0, 1, 2, \dots, K - 1,$$

and

$$\int_{(-\infty, t]} |x|^{K+1} \mu(dx) \leq C(\varepsilon)t^{-\alpha+K+1+\varepsilon}$$

for any $t \geq 1$.

The following is well known.

PROPOSITION 9. (1) *For any $m \geq 0$, let $r_{e,m} : \mathbf{R} \rightarrow \mathbf{C}$ be given by*

$$r_{e,m}(t) = \exp(it) - \left(1 + \sum_{k=1}^m \frac{(it)^k}{k!}\right), \quad t \in \mathbf{R}.$$

Then we have

$$|r_{e,m}(t)| \leq \frac{|t|^{m+1}}{(m+1)!} \quad t \in \mathbf{R}.$$

(2) For any $m \geq 1$, let $r_{l,m} : \{z \in \mathbf{C}; |z| \leq 1/2\} \rightarrow \mathbf{C}$ be given by

$$r_{l,m}(z) = \log(1+z) - \sum_{k=1}^m \frac{(-1)^{k-1}}{k} z^k, \quad z \in \mathbf{C}, |z| \leq 1/2.$$

Then we have

$$|r_{l,m}(z)| \leq 2|z|^{m+1}, \quad z \in \mathbf{C}, |z| \leq 1/2.$$

Let $\mu(t)$, $t > 0$, be a probability measure on $(\mathbf{R}, \mathcal{B}(\mathbf{R}))$ given by

$$\mu(t)(A) = (1 - \bar{F}(t))^{-1} \mu(A \cap (-\infty, t]),$$

for any $A \in \mathcal{B}(\mathbf{R})$.

Let $\varphi(\cdot; \mu(t))$, $t > 0$, be the characteristic function of the probability measure $\mu(t)$, i.e.,

$$\varphi(\xi; \mu(t)) = \int_{\mathbf{R}} \exp(ix\xi) \mu(t)(dx), \quad \xi \in \mathbf{R}.$$

By the assumption (A3), we see that the density function $\rho(x) \rightarrow 0$ as $|x| \rightarrow \infty$. Also we see that the probability measure $\mu(t)$, $t \geq 2$, is absolutely continuous and its density function is $(1 - \bar{F}(t))^{-1} \rho(x) 1_{(-\infty, t]}(x)$, whose total variation is dominated by twice of that of ρ .

Therefore we have the following.

PROPOSITION 10. (1) For any $t \geq 2$ and $\xi \in \mathbf{R}$,

$$\begin{aligned} i\xi \varphi(\xi; \mu(t)) &= (1 - \bar{F}(t))^{-1} \int_{\mathbf{R}} i\xi e^{i\xi x} \rho(x) 1_{(-\infty, t]}(x) dx \\ &= -(1 - \bar{F}(t))^{-1} \int_{\mathbf{R}} e^{i\xi x} d(\rho(x) 1_{(-\infty, t]}(x)). \end{aligned}$$

(2) There is a $C > 0$ such that

$$|\varphi(\xi, \mu(t))| \leq C(1 + |\xi|)^{-1} \text{ for any } t \geq 2 \text{ and } \xi \in \mathbf{R}.$$

Then we have the following.

PROPOSITION 11. (1) *There is a $c_0 > 0$ such that*

$$|\varphi(\xi, \mu(t))| \leq (1 + c_0|\xi|^2)^{-1/4} \text{ for any } t \geq 2 \text{ and } \xi \in \mathbf{R}.$$

(2) *For any $t \geq 2$, $\xi \in \mathbf{R}$, and integers n, m with $n \geq m$,*

$$|\varphi(n^{-1/2}\xi, \mu(t))|^n \leq (1 + \frac{c_0}{m}|\xi|^2)^{-m/4}.$$

PROOF. Let $g(x) = \rho(x)1_{(-2,2)}(x)$, $x \in \mathbf{R}$. Then we have

$$p = \int_{\mathbf{R}} g(x)dx \geq 1 - \int_{\mathbf{R}} \frac{x^2}{4}\rho(x)dx \geq 3/4.$$

Note that

$$\begin{aligned} & |\varphi(\xi, \mu(t))|^2 \\ &= (1 - \bar{F}(t))^{-2} \int_{\mathbf{R}} \int_{\mathbf{R}} \exp(i\xi(x-y))\rho(x)1_{(-\infty,t]}(x)\rho(y)1_{(-\infty,t]}(y)dxdy \\ &\leq (1 - p^2) + \int_{\mathbf{R}} \int_{\mathbf{R}} \exp(i\xi(x-y))g(x)g(y)dxdy = 1 - f(\xi), \end{aligned}$$

where

$$f(\xi) = \int_{\mathbf{R}} \int_{\mathbf{R}} (1 - \cos(\xi(x-y)))g(x)g(y)dxdy.$$

So we see that

$$\lim_{\xi \rightarrow 0} |\xi|^{-2}f(\xi) = \frac{1}{2} \int_{\mathbf{R}} \int_{\mathbf{R}} (x-y)^2g(x)g(y)dxdy > 0.$$

Also, it is easy to see that $f(\xi) > 0$, for all $\xi \in \mathbf{R} \setminus \{0\}$, and so we see that

$$a(r) = \inf_{|\xi| \leq r} |\xi|^{-2}f(\xi) > 0 \quad \text{for all } r > 0.$$

Therefore we see that

$$|\varphi(\xi, \mu(t))| \leq (1 - a(r)|\xi|^2)^{1/2} \leq (1 + a(r)|\xi|^2)^{-1/4}, \quad |\xi| \leq r.$$

Also by Proposition 10(2), we see that there is an $r_0 > 0$ such that

$$|\varphi(\xi, \mu(t))| \leq (1 + |\xi|^2)^{-1/4}, \quad |\xi| \geq r_0$$

So we have the assertion (1).

It is easy to check that $(1 + x/\beta)^\beta \geq 1 + x$ for any $\beta \geq 1$ and $x \geq 0$. Therefore if $n \geq m$, we have

$$(1 + c_0|n^{-1/2}\xi|^2)^{n/m} \geq 1 + \frac{c_0}{m}|\xi|^2.$$

This implies the assertion (2). \square

5. Asymptotic Expansion of Characteristic Functions

Let

$$\varphi_1(\xi, t) = - \sum_{k=1}^{K-1} \frac{(i\xi)^k}{k!} \bar{\eta}_k(t) + \frac{(i\xi)^K}{K!} \eta_K(t)$$

and

$$\begin{aligned} \psi_0(n, \xi) &= \sum_{k=3}^{K-1} n^{-k/6} q_k(n^{-1/3}, \eta_2, \dots, \eta_k) (i\xi)^k \\ &\quad + \sum_{k=K}^{3(K-1)} n^{-k/6} q_k(n^{-1/3}, \eta_2, \dots, \eta_{K-1}, 0, \dots, 0) (i\xi)^k \end{aligned}$$

for $t \geq 2$, $n \geq 1$ and $\xi \in \mathbf{R}$. Let $\delta = ((\alpha-2) \wedge 1)/(4(K+2))$, $\delta' = \delta/(4(K+2))$, and $t_n = n^{1/2-\delta}$, $n = 1, 2, 3, \dots$. Then $t_n \geq 2$ for any $n \geq 8$.

In this section, we prove the following.

LEMMA 12. *Let*

$$R_{n,0}(\xi) = \exp\left(\frac{1}{2}\xi^2\right)\varphi(n^{-1/2}\xi, \mu(t_n))^n - (1 + \psi_0(n, \xi) + n\varphi_1(n^{-1/2}\xi, t_n))$$

$$R_{n,1}(\xi) = \exp\left(\frac{1}{2}\xi^2\right)\varphi(n^{-1/2}\xi, \mu(t_n))^n - 1$$

$$R_{n,2}(\xi) = \exp\left(\frac{1}{2}\xi^2\right)\varphi(n^{-1/2}\xi, \mu(t_n))^{n-1} - 1$$

Then there is a $C > 0$ such that

$$|R_{n,0}(\xi)| \leq Cn^{-(\alpha-2)/2-\delta/4}|\xi|$$

and

$$|R_{n,1}(\xi)| + |R_{n,2}(\xi)| \leq Cn^{-2K\delta}|\xi|$$

for any $n \geq 8$ and $\xi \in \mathbf{R}$ with $|\xi| \leq n^{\delta'}$.

We make some preparations to prove this lemma. First we prove the following.

PROPOSITION 13. *Let*

$$\varphi_0(\xi) = \sum_{k=2}^{K-1} \frac{(i\xi)^k}{k!} \eta_k,$$

and

$$R_0(\xi, t) = \varphi(\xi; \mu(t)) - (1 + \varphi_0(\xi) + \varphi_1(\xi, t)).$$

Then we have for any $n \geq 8$, and $\xi \in \mathbf{R}$ with $|\xi| \leq n^{\delta'}$,

$$|\varphi(n^{-1/2}\xi; \mu(t_n)) - 1| \leq \frac{2\sqrt{3}}{3}n^{-1/2}|\xi|,$$

$$|\varphi_1(n^{-1/2}\xi, t_n)| \leq KC(\delta)n^{-\alpha/2+(K+1)\delta}|\xi|$$

and

$$|R_0(n^{-1/2}\xi, t_n)| \leq 3C(\delta)n^{-\alpha/2-\delta/4}|\xi|.$$

Here $C(\delta)$ is as in Proposition 8.

PROOF. We can easily see that

$$\begin{aligned} \varphi(\xi; \mu(t)) &= \int_{\mathbf{R}} \exp(ix\xi)\mu(t)(dx) \\ &= 1 + \sum_{k=1}^K \frac{(i\xi)^k}{k!} \eta_k(t) + \int_{(-\infty, t]} r_{e,K}(x\xi)\mu(dx) \\ &\quad + \bar{F}(t)(1 - \bar{F}(t))^{-1} \int_{(-\infty, t]} r_{e,0}(x\xi)\mu(dx) \end{aligned}$$

So we see that

$$R_0(\xi, t) = \bar{F}(t)(1 - \bar{F}(t))^{-1} \int_{(-\infty, t]} r_{e,0}(x\xi)\mu(dx) + \int_{(-\infty, t]} r_{e,K}(x\xi)\mu(dx).$$

By Propositions 8 and 9 we have

$$|\varphi_1(\xi, t)| \leq C(\delta) \sum_{k=1}^K \frac{|\xi|^k}{k!} t^{-\alpha+k+\delta}, \quad \xi \in \mathbf{R}, t \geq 2,$$

and

$$|R_0(\xi, t)| \leq C(\delta)|\xi|t^{-\alpha+\delta} \int_{\mathbf{R}} |x|\mu(t)(dx) + C(\delta)|\xi|^{K+1}t^{-\alpha+K+1+\delta},$$

$$\xi \in \mathbf{R}, t \geq 2.$$

Also, we have

$$|\varphi(\xi; \mu(t)) - 1| \leq |\xi| \int_{\mathbf{R}} |x|\mu(t)(dx) \leq (1 - \bar{F}(t))^{-1/2}|\xi| \leq \frac{2\sqrt{3}}{3}|\xi|,$$

$$\xi \in \mathbf{R}, t \geq 2.$$

Note that

$$(n^{-1/2+\delta'})^k (n^{1/2-\delta})^{-\alpha+k+\delta} = n^{-\alpha/2+(\alpha+1/2)\delta-k(\delta-\delta')-\delta^2}.$$

So we have our assertion. \square

PROPOSITION 14. *Let*

$$\psi_1(\xi) = \sum_{k=3}^{K-1} \frac{(i\xi)^k}{k!} c_k(\eta_2, \dots, \eta_k) + \frac{(i\xi)^K}{K!} c_K(\eta_2, \dots, \eta_{K-1}, 0), \quad \xi \in \mathbf{R}.$$

Also, for any $n \geq 8$, and $\xi \in \mathbf{R}$ with $|\xi| \leq n^{\delta'}$, let

$$R_1(n, \xi) = \log(\varphi(n^{-1/2}\xi, \mu(t_n))) - \left\{ -\frac{1}{2n}\xi^2 + \psi_1(n^{-1/2}\xi) + \varphi_1(n^{-1/2}\xi, t_n) \right\}.$$

Then there is a constant $C > 0$ such that

$$|R_1(n, \xi)| \leq Cn^{-\alpha/2-\delta/4}|\xi|$$

for any $n \geq 8$, and $\xi \in \mathbf{R}$ with $|\xi| \leq n^{\delta'}$.

PROOF. Let

$$R_{1,1}(\xi) = \sum_{k=1}^K \frac{(-1)^{k-1}}{k} (\varphi_0(\xi))^k + \frac{1}{2} \xi^2 - \psi_1(\xi).$$

Note that

$$\begin{aligned} & \log\left(1 + \sum_{k=2}^{K-1} \eta_k \frac{z^k}{k!}\right) \\ &= \sum_{k=2}^{K-1} c_k(\eta_2, \dots, \eta_k) \frac{z^k}{k!} + \sum_{k=K}^{\infty} c_k(\eta_2, \dots, \eta_{K-1}, 0, \dots, 0) \frac{z^k}{k!} \end{aligned}$$

as a formal power series of z . So we see that there is a constant $C > 0$ such that

$$(4) \quad |R_{1,1}(\xi)| \leq C|\xi|^{K+1}$$

for any $\xi \in \mathbf{R}$ with $|\xi| \leq 1$.

We can easily see that

$$\begin{aligned} & R_1(n, \xi) \\ &= \log(1 + \varphi_0(n^{-1/2}\xi) + \varphi_1(n^{-1/2}\xi, t_n) + R_0(n^{-1/2}\xi, t_n)) \\ & \quad - \left\{ -\frac{1}{2n} \xi^2 + \psi_1(n^{-1/2}\xi) + \varphi_1(n^{-1/2}\xi, t_n) \right\} \\ &= R_{1,1}(n^{-1/2}\xi) + r_{l,K}(\varphi(n^{-1/2}\xi, \mu(t_n)) - 1) + R_0(n^{-1/2}\xi, t_n) \\ & \quad + \sum_{k=2}^K (-1)^{k-1} (\varphi_0(n^{-1/2}\xi))^{k-1} (\varphi_1(n^{-1/2}\xi, t_n) + R_0(n^{-1/2}\xi, t_n)) \\ & \quad + \sum_{k=1}^K \frac{(-1)^{k-1}}{k} \sum_{j=2}^k \binom{k}{j} (\varphi_0(n^{-1/2}\xi))^{k-j} (\varphi_1(n^{-1/2}\xi, t_n) + R_0(n^{-1/2}\xi, t_n))^j. \end{aligned}$$

Then we have our assertion from Equation (4) and Proposition 13.

PROPOSITION 15. *Let*

$$R_2(n, \xi) = \exp(n\psi_1(n^{-1/2}\xi)) - (1 + \psi_0(n, \xi)).$$

Then there is a constant $C > 0$ such that

$$|R_2(n, \xi)| \leq Cn^{-(\alpha-2)/2-1/4}|\xi|$$

for any $n \geq 8$, and $\xi \in \mathbf{R}$ with $|\xi| \leq n^{\delta'}$.

PROOF. Note that

$$\begin{aligned} & \exp(y^{-6}(\sum_{k=3}^{K-1} \frac{(y^3 z)^k}{k!} c_k(\eta_2, \dots, \eta_k) + \sum_{k=K}^{\infty} \frac{(y^3 z)^k}{k!} c_k(\eta_2, \dots, \eta_{K-1}, 0, \dots, 0))) \\ &= 1 + \sum_{k=3}^{K-1} y^k q_k(y^2, \eta_2, \dots, \eta_k) z^k + \sum_{k=K}^{\infty} y^k q_k(\eta_2, \dots, \eta_{K-1}, 0, \dots, 0) z^k \end{aligned}$$

as a formal power series in z . This implies our assertion. \square

Now let us prove Lemma 12.

Note that for any $n \geq 8$, and $\xi \in \mathbf{R}$ with $|\xi| \leq n^{\delta'}$,

$$\begin{aligned} & \exp(\frac{1}{2}\xi^2)\varphi(n^{-1/2}\xi; \mu(t_n))^n \\ &= \exp(n\varphi_1(n^{-1/2}\xi, t_n) + n\psi_1(n^{-1/2}\xi) + nR_1(n, \xi)) \\ &= (1 + n\varphi_1(n^{-1/2}\xi, t_n) + r_{e,1}(n\varphi_1(n^{-1/2}\xi, t_n))) \\ & \quad \times (1 + \psi_0(n, \xi) + R_2(n, \xi))(1 + r_{e,0}(nR_1(n, \xi))). \end{aligned}$$

So we see that

$$\begin{aligned} R_{n,0}(n, \xi) &= r_{e,0}(nR_1(n, \xi)) \exp(n\varphi_1(n^{-1/2}\xi, t_n) + n\psi_1(n^{-1/2}\xi)) \\ & \quad + R_2(n, \xi) \exp(n\varphi_1(n^{-1/2}\xi, t_n)) \\ & + r_{e,1}(n\varphi_1(n^{-1/2}\xi, t_n)) + \psi_0(n, \xi)(n\varphi_1(n^{-1/2}\xi, t_n) + r_{e,1}(n\varphi_1(n^{-1/2}\xi, t_n))). \end{aligned}$$

Thus we have the first equation from Propositions 13, 14, 15.

Also, we have

$$R_{n,1}(n, \xi) = \exp(n\varphi_1(n^{-1/2}\xi, t_n) + n\psi_1(n^{-1/2}\xi) + nR_1(n, \xi)) - 1,$$

and

$$\begin{aligned} R_{n,2}(n, \xi) &= \exp((n-1)\varphi_1(n^{-1/2}\xi, t_n) + (n-1)\psi_1(n^{-1/2}\xi) \\ & \quad + (n-1)R_1(n, \xi)) - \frac{\xi^2}{n} - 1. \end{aligned}$$

So, again from Propositions 13, 14, 15 we have the second equation.

6. Proof of Theorem 2

First, we prove the following.

LEMMA 16. *Let ν be a probability measure on $(\mathbf{R}, \mathcal{B}(\mathbf{R}))$ such that $\int_{\mathbf{R}} x^2 \nu(dx) < \infty$. Also, assume that there is a constant $C > 0$ such that the characteristic function $\varphi(\cdot, \nu) : \mathbf{R} \rightarrow \mathbf{C}$ satisfies*

$$|\varphi(\xi; \nu)| \leq C(1 + |\xi|)^{-2}, \quad \xi \in \mathbf{R}.$$

Then for any $x \in \mathbf{R}$

$$\nu((x, \infty)) = \Phi_0(x) + \frac{1}{2\pi} \int_{\mathbf{R}} \frac{e^{-ix\xi}}{i\xi} (\varphi(\xi, \nu) - \exp(-\frac{\xi^2}{2})) d\xi.$$

PROOF. From the assumption, ν has a continuous density function β and

$$\beta(x) = \frac{1}{2\pi} \int_{\mathbf{R}} e^{-ix\xi} \varphi(\xi, \nu) d\xi.$$

So we have

$$\begin{aligned} \nu((x, x+n]) &= \Phi_0(x) - \Phi_0(x+n) \\ &+ \frac{1}{2\pi} \int_{\mathbf{R}} \left(\int_x^{x+n} e^{-iz\xi} dz \right) (\varphi(\xi, \nu) - \exp(-\frac{\xi^2}{2})) d\xi. \\ &= \Phi_0(x) - \Phi_0(x+n) + \frac{1}{2\pi} \int_{\mathbf{R}} \frac{e^{-ix\xi} - e^{-i(x+n)\xi}}{i\xi} (\varphi(\xi, \nu) - \exp(-\frac{\xi^2}{2})) d\xi. \end{aligned}$$

Since

$$\int_{\mathbf{R}} \frac{1}{|\xi|} |\varphi(\xi, \nu) - \exp(-\frac{\xi^2}{2})| d\xi < \infty,$$

letting $n \rightarrow \infty$, we have the assertion. \square

Note that

$$P\left(\sum_{k=1}^n X_k > sn^{1/2}\right) = \sum_{m=0}^n I_m(n, s),$$

where

$$I_m(n, s) = P\left(\sum_{k=1}^n X_k > sn^{1/2}, \sum_{k=1}^n 1_{\{X_k > t_n\}} = m\right), \quad m = 0, 1, \dots, n.$$

Then we have

$$I_m(n, s) = \binom{n}{m} P\left(\sum_{k=1}^n X_k > sn^{1/2}, X_i > t_n, i = 1, \dots, m, \right. \\ \left. X_j \leq t_n, j = m + 1, \dots, n\right),$$

for $m = 0, 1, \dots, n$.

PROPOSITION 17. *There is a $C > 0$ such that*

$$\sum_{m=2}^n I_m(n, s) \leq Cn^{-(\alpha-2)/2-\delta}$$

for any $s \geq 1$ and $n \geq 8$.

PROOF. We see that by Proposition 8

$$\sum_{m=2}^n I_m(n, s) \leq \sum_{m=2}^n \frac{n(n-1)}{m(m-1)} \binom{n-2}{m-2} \bar{F}(t_n)^m (1 - \bar{F}(t_n))^{n-m} \\ \leq \frac{n(n-1)}{2} \bar{F}(t_n)^2 \leq C(\delta)^2 n^{-(\alpha-2)/2-\delta}.$$

This implies our assertion. \square

PROPOSITION 18. *There is a $C > 0$ such that*

$$\sup_{s \in [1, \log n]} |I_0(n, s) - \{(1 - n\bar{F}(t_n))\Phi_0(s) - \sum_{k=1}^{K-1} \frac{n^{-(k-2)/2}}{k!} \bar{\eta}_k(t_n) \Phi_k(s) \\ + \frac{(n^{1/2})^{K-2}}{K!} \eta_K(t_n) \Phi_K(s) + g(n, s)\}| \leq Cn^{-(\alpha-2)/2-\delta/4}$$

for any $n \geq 8$. Here

$$g(n, s) = \sum_{k=3}^{K-1} n^{-k/6} q_k(n^{-1/3}, \eta_2, \dots, \eta_k) \Phi_k(s) \\ + \sum_{k=K}^{3(K-1)} n^{-k/6} q_k(n^{-1/3}, \eta_2, \dots, \eta_{K-1}, 0, \dots, 0) \Phi_k(s).$$

PROOF. Note that

$$\begin{aligned} I_0(n, s) &= (1 - \bar{F}(t_n))^n \mu(t_n)^{*n}((sn^{1/2}, \infty)) \\ &= I_{0,0}(n, s) + I_{0,1}(n, s) + I_{0,2}(n, s), \end{aligned}$$

where

$$\begin{aligned} I_{0,0}(n, s) &= \mu(t_n)^{*n}((sn^{1/2}, \infty)), \\ I_{0,1}(n, s) &= -n\bar{F}(t_n)\mu(t_n)^{*n}((sn^{1/2}, \infty)), \\ I_{0,2}(n, s) &= ((1 - \bar{F}(t_n))^n - 1 + n\bar{F}(t_n))\mu(t_n)^{*n}((sn^{1/2}, \infty)). \end{aligned}$$

We remark that

$$\Phi_k(x) = \frac{1}{2\pi} \int_{-\infty}^{\infty} (i\xi)^{k-1} \exp(-i\xi x - \frac{\xi^2}{2}) d\xi, \quad k = 1, 2, \dots$$

By Proposition 11 and Lemma 16, we have

$$\begin{aligned} &I_{0,0}(n, s) \\ &= \Phi_0(s) + \frac{1}{2\pi} \int_{\mathbf{R}} \frac{e^{-is\xi}}{i\xi} (\varphi(n^{-1/2}\xi, \mu(t_n))^n - \exp(-\frac{\xi^2}{2})) d\xi. \end{aligned}$$

Let

$$\begin{aligned} &\tilde{R}_{0,0}(n, s) = I_{0,0}(n, s) \\ &- \left\{ \Phi_0(s) + \frac{1}{2\pi} \int_{\mathbf{R}} \frac{e^{-is\xi}}{i\xi} (\psi_0(n, \xi) + n\varphi_1(n^{-1/2}\xi, t_n)) e^{-\xi^2/2} d\xi \right\} \end{aligned}$$

Then by Lemma 12 we have

$$\begin{aligned} &|\tilde{R}_{0,0}(n, s)| \\ &\leq \int_{|\xi| \leq n^{\delta'}} \frac{|R_{n,0}(\xi)|}{|\xi|} \exp(-\frac{\xi^2}{2}) d\xi \\ &+ \int_{|\xi| > n^{\delta'}} \frac{1}{|\xi|} (|\varphi(n^{-1/2}\xi, \mu(t_n))|^n + \exp(-\frac{\xi^2}{2})) d\xi \\ &+ \int_{|\xi| > n^{\delta'}} (|\psi_0(n, \xi)| + n|\varphi_1(n^{-1/2}\xi, t_n)|) e^{-\xi^2/2} d\xi \end{aligned}$$

So by Proposition 11 and Lemma 12, we see that there is a $C_0 > 0$ such that

$$(5) \quad |\tilde{R}_{0,0}(n, s)| \leq C_0 n^{-(\alpha-2)/2-\delta/4}, \quad n \geq 8, s \geq 1.$$

Also, we see that

$$\begin{aligned} & \frac{1}{2\pi} \int_{\mathbf{R}} \frac{e^{-is\xi}}{i\xi} n \varphi_1(n^{-1/2}\xi, t_n) \exp(-\frac{1}{2}\xi^2) d\xi \\ &= - \sum_{k=1}^{K-1} \frac{(n^{-1/2})^{k-2}}{k!} \bar{\eta}_k(t_n) \Phi_k(s) + \frac{(n^{-1/2})^{K-2}}{K!} \eta_K(t_n) \Phi_K(s), \end{aligned}$$

and

$$\frac{1}{2\pi} \int_{\mathbf{R}} \frac{e^{-is\xi}}{i\xi} \psi_0(n, \xi) \exp(-\frac{1}{2}\xi^2) d\xi = g(n, s)$$

Similarly by Lemma 12, we see that there is a $C_1 > 0$ such that

$$(6) \quad \sup_{s \in [1, \log n]} |I_{0,1}(n, s) - n \bar{F}(t_n) \Phi_0(s)| \leq C_1 n^{-(\alpha-2)/2-\delta}, \quad n \geq 8.$$

Note that $|(1-x)^n - (1-nx)| \leq n^2 x^2$ for any $x \in [0, 1]$, $n \geq 1$. So we have

$$|I_{0,2}(n, s)| \leq n^2 \bar{F}(t_n)^2 \leq C(\delta)^2 n^{-(\alpha-2)/2-\delta}.$$

This and Equations 5, 6 imply our assertion. \square

PROPOSITION 19. *There is a $C > 0$ such that*

$$\begin{aligned} & \sup_{s \in [1, \log n]} |I_1(n, s) - \{n \int_{-\infty}^s \bar{F}((s-x)n^{1/2}) \Phi_1(x) dx + n \bar{F}(t_n) \Phi_0(s) \\ & \quad - \sum_{k=1}^K \frac{n^{-(k-2)/2}}{k!} \Phi_k(s) \int_0^{t_n} x^k \mu(dx)\}| \leq C n^{-(\alpha-2)/2-\delta/4}. \end{aligned}$$

PROOF. We see that

$$I_1(n, s) = n(1 - \bar{F}(t_n))^{n-1} \int_{\mathbf{R}} P(X_1 + x > sn^{1/2}, X_1 > t_n) \mu(t_n)^{*(n-1)}(dx)$$

$$\begin{aligned}
&= n(1 - \bar{F}(t_n))^{n-1} \int_{-\infty}^{\infty} \bar{F}((sn^{1/2} - x) \vee t_n) \mu(t_n)^{* (n-1)}(dx) \\
&= nJ_0(n, s) + nJ_1(n, s) + nJ_2(n, s),
\end{aligned}$$

where

$$(7) \quad J_0(n, s) = \int_{-\infty}^{\infty} \bar{F}((s - x)n^{1/2} \vee t_n) \Phi_1(x) dx,$$

$$(8) \quad \begin{aligned} &J_1(n, s) \\ &= \int_{-\infty}^{\infty} \bar{F}((sn^{1/2} - x) \vee t_n) (\mu(t_n)^{* (n-1)}(dx) - n^{-1/2} \Phi_1(xn^{1/2}) dx), \end{aligned}$$

and

$$(9) \quad J_2(n, s) = -(1 - (1 - \bar{F}(t_n))^{n-1}) I_1(n, s).$$

Note that

$$J_0(n, s) = J_{0,0}(n, s) + J_{0,1}(n, s) + J_{0,2}(n, s),$$

where

$$\begin{aligned}
J_{0,0}(n, s) &= \int_{-\infty}^s \bar{F}((s - x)n^{1/2}) \Phi_1(x) dx, \\
J_{0,1}(n, s) &= - \int_{s-n^{-\delta}}^s \bar{F}((s - x)n^{1/2}) \Phi_1(x) dx \\
&= - \int_0^{n^{-\delta}} \bar{F}(xn^{1/2}) \Phi_1(s - x) dx,
\end{aligned}$$

and

$$J_{0,2}(n, s) = \bar{F}(t_n) \int_{s-n^{-\delta}}^{\infty} \Phi_1(x) dx = \bar{F}(t_n) \Phi_0(s - n^{-\delta}).$$

We see that

$$J_{0,1}(n, s) = - \sum_{k=1}^K \frac{1}{(k-1)!} \Phi_k(s) \int_0^{n^{-\delta}} \bar{F}(xn^{1/2}) x^{k-1} dx + R_{J,1}(n, s),$$

where

$$R_{J,1}(n, s) = - \int_0^{n^{-\delta}} \bar{F}(xn^{1/2}) (\Phi_1(s - x) - \sum_{k=1}^K \frac{x^{k-1}}{(k-1)!} \Phi_k(s)) dx.$$

Then

$$\begin{aligned}
& |R_{J,1}(n, s)| \\
\leq & \sup_{x \in [0, n^{-\delta}]} |\Phi_{K+1}(s-x)| \left(\int_{n^{-1/2}}^{n^{-\delta}} x^K (xn^{1/2})^{-\alpha} L(xn^{1/2}) dx + \int_0^{n^{-1/2}} x^K dx \right) \\
\leq & \sup_{x \in \mathbf{R}} |\Phi_{K+1}(x)| \left(C(\delta) n^{-\alpha/2+\delta/2} \int_0^{n^{-\delta}} x^{\delta+(K-\alpha)} dx + n^{-(K+1)/2} \right) \\
(10) \quad & \leq \sup_{x \in \mathbf{R}} |\Phi_{K+1}(x)| (C(\delta) + 1) n^{-\alpha/2-\delta/2}.
\end{aligned}$$

Also, we see that

$$\begin{aligned}
& J_{0,2}(n, s) \\
= & \bar{F}(t_n) \Phi_0(s) + \sum_{k=1}^K \bar{F}(t_n) \frac{(n^{-\delta})^k}{k!} \Phi_k(s) + R_{J,2}(n, s),
\end{aligned}$$

where

$$R_{J,2}(n, s) = \bar{F}(t_n) (\Phi_0(s - n^{-\delta}) - \sum_{k=0}^K \frac{(-n^{-\delta})^k}{k!} \frac{d^k \Phi_0}{dx^k}(s)).$$

We see that

$$\begin{aligned}
(11) \quad |R_{J,2}(n, s)| & \leq \bar{F}(t_n) n^{-(K+1)\delta} \sup_{x \in \mathbf{R}} |\Phi_{K+1}(x)| \\
& \leq C(\delta) \sup_{x \in \mathbf{R}} |\Phi_{K+1}(x)| n^{-\alpha/2-\delta/4}.
\end{aligned}$$

It is easy to see that

$$\begin{aligned}
& \int_0^{n^{-\delta}} \bar{F}(xn^{1/2}) x^{k-1} dx = n^{-k/2} \int_0^{t_n} \bar{F}(x) x^{k-1} dx \\
= & n^{-k/2} \left(-\frac{1}{k} \int_0^{t_n} x^k \mu(dx) + \frac{n^{\delta k}}{k} \bar{F}(t_n) \right), \quad k = 1, \dots, K.
\end{aligned}$$

So we have

$$J_{0,1}(n, s) + J_{0,2}(n, s) = \bar{F}(t_n) \Phi_0(s)$$

$$(12) \quad - \sum_{k=1}^K \frac{n^{-k/2}}{k!} \Phi_k(s) \int_0^{t_n} x^k \mu(dx) + R_{J,1}(n, s) + R_{J,2}(n, s)$$

Also, we have

$$J_1(n, s) = J_{1,1}(n, s) + J_{1,2}(n, s)$$

where

$$J_{1,1}(n, s) = \bar{F}(t_n)(\mu(t_n)^{*n-1}((s - n^{-\delta})n^{1/2}, \infty)) - \Phi_0(s - n^{-\delta})$$

and

$$\begin{aligned} J_{1,2}(n, s) &= \int_{-\infty}^{s-n^{-\delta}} dx \bar{F}((s-x)n^{1/2}) \\ &\times \frac{1}{2\pi} \int_{\mathbf{R}} e^{-ix\xi} (\varphi(n^{-1/2}\xi; \mu(t_n))^{n-1} - \exp(-\frac{\xi^2}{2})) d\xi \end{aligned}$$

By Proposition 11 and Lemma 16, we see that there is a $C_1 > 0$ such that

$$\begin{aligned} &|\mu(t_n)^{*n-1}((xn^{1/2}, \infty)) - \Phi_0(x)| \\ &\leq \left| \int_{-\infty}^{\infty} \frac{e^{-ix\xi}}{\xi} (\varphi(n^{-1/2}\xi; \mu(t_n))^{n-1} - \exp(-\frac{\xi^2}{2})) d\xi \right| \\ &\leq \int_{|\xi| > n^{\delta'}} \frac{1}{|\xi|} (|\varphi(\xi; \mu(t_n))|^{n-1} + \exp(-\frac{\xi^2}{2})) d\xi \\ &\quad + \int_{|\xi| < n^{\delta'}} \frac{1}{|\xi|} |R_{n,2}(\xi)| \exp(-\frac{\xi^2}{2}) d\xi \\ &\leq C_1 n^{-2K\delta}, \text{ for any } x \in \mathbf{R} \text{ and } n \geq 8. \end{aligned}$$

Therefore we have

$$|J_{1,1}(n, s)| \leq C_1 \bar{F}(t_n) n^{-2K\delta} \leq C(\delta) C_1 n^{-\alpha/2-\delta}.$$

Similarly by Lemma 16, we see that there is a $C_2 > 0$ such that

$$\begin{aligned} &\left| \int_{\mathbf{R}} e^{-ix\xi} (\varphi(n^{-1/2}\xi; \mu(t_n))^{n-1} - \exp(-\frac{\xi^2}{2})) d\xi \right| \\ &\leq C_2 n^{-2K\delta}, \text{ for any } x \in \mathbf{R} \text{ and } n \geq 8. \end{aligned}$$

Then we have

$$|J_{1,2}(n, s)| \leq C_2 n^{-2K\delta} C(\delta) \int_{n^{-\delta}}^{\infty} (xn^{1/2})^{-\alpha+\delta} dx \leq C_2 C(\delta) n^{-\alpha/2-\delta}$$

So we see that there is a $C > 0$ such that

$$(13) \quad \sup_{s \in [1, \log n]} |J_1(n, s)| \leq C n^{-(\alpha-2)/2-\delta}$$

Note that

$$(14) \quad |J_2(n, s)| \leq n^2 \bar{F}(t_n)^2$$

So Equations (7) - (14) imply our assertion. \square

PROPOSITION 20. *Then there is a $C > 0$ such that*

$$\sup_{s \in [1, \log n]} |P(\sum_{k=1}^n X_k > sn^{1/2}) - G(n, s)| \leq C n^{-(\alpha-2)/2-\delta/4}, \quad n = 3, 4, \dots$$

PROOF. Note that

$$\bar{\eta}_k(t_n) + \int_0^{t_n} x^k \mu(dx) = \int_0^{\infty} x^k \mu(dx), \quad k = 1, 2, \dots, K-1,$$

and

$$\eta_K(t_n) - \int_0^{t_n} x^K \mu(dx) = \int_{-\infty}^0 x^K \mu(dx).$$

So our assertion is an easy consequence of Propositions 17, 18, 19. \square

PROPOSITION 21. *There is a $C > 0$ such that*

$$\sup_{s \in [\log n, \infty)} I_0(n, s) \leq C n^{-(\alpha-2)/2-\delta/4}, \quad n = 3, 4, \dots$$

PROOF. We have

$$I_0(n, s) \leq \mu(t_n)^{*n}((sn^{1/2}, \infty))$$

$$\begin{aligned}
&= \Phi_0(s) + \frac{1}{2\pi} \int_{\mathbf{R}} \frac{e^{-is\xi}}{i\xi} (\varphi(n^{-1/2}\xi, \mu(t_n))^n - \exp(-\frac{\xi^2}{2})) d\xi. \\
&= \Phi_0(s) + \frac{1}{2\pi} \int_{\mathbf{R}} \frac{e^{-is\xi}}{i\xi} (\varphi(n^{-1/2}\xi, \mu(t_n))^n - \exp(-\frac{\xi^2}{2})) d\xi. \\
&\leq \Phi_0(s) + \frac{1}{2\pi} \int_{\mathbf{R}} \frac{e^{-is\xi}}{i\xi} (\psi_0(n, \xi) + n\varphi_1(n^{-1/2}\xi, t_n) \exp(-\frac{\xi^2}{2})) d\xi. \\
&\quad + \frac{1}{2\pi} \int_{\mathbf{R}} \frac{e^{-is\xi}}{i\xi} R_{n,0}(\xi) \exp(-\frac{\xi^2}{2}) d\xi.
\end{aligned}$$

Since $\sup_{s \geq \log n} |\Phi_k(s)|$ is of $O(n^{-M})$ for any $M \geq 1$, we have our assertion similar to the proof of Proposition 18.

PROPOSITION 22. *There is a $C > 0$ such that*

$$\sup_{s \in [\log n, \infty)} |I_1(n, s) - n \int_{-\infty}^s \bar{F}((s-x)n^{1/2}) \Phi_1(x) dx| \leq Cn^{-(\alpha-2)/2-\delta/4}.$$

PROOF. Remind that

$$\begin{aligned}
I_1(n, s) &= n(1 - \bar{F}(t_n))^{n-1} \int_{-\infty}^{\infty} \bar{F}((sn^{1/2} - x) \vee t_n) \mu(t_n)^{*n-1}(dx) \\
&= n(1 - \bar{F}(t_n))^{n-1} \left\{ \int_{-\infty}^{\infty} \bar{F}(((s-x)n^{1/2}) \vee t_n) \Phi_1(x) dx \right. \\
&\quad \left. + \frac{1}{2\pi} \int_{-\infty}^{\infty} \bar{F}(((s-x)n^{1/2}) \vee t_n) \right. \\
&\quad \left. \times \left(\int_{\mathbf{R}} \frac{e^{-ix\xi}}{i\xi} (\varphi(n^{-1/2}\xi, \mu(t_n))^{n-1} - \exp(-\frac{\xi^2}{2})) d\xi \right) \right\}.
\end{aligned}$$

Then similarly to the proof of Propositions 19 and 21, we have our assertion. \square

Now Theorem 2 is a consequence of Propositions 20 and 22.

7. Preliminary for Theorem 3

PROPOSITION 23. *Let Y be a random variable, and assume that*

$$E[|Y|^2] < \infty \quad \text{and} \quad E[Y] = 0.$$

Then for any $s \in \mathbf{R} \setminus \{0\}$ and $b > 0$

$$E[\exp(sY1_{\{|Y| \leq b\}})] \leq 1 + |s|^2 \left(1 + \frac{1}{|s|b}\right) \exp(|s|b) E[|Y|^2].$$

PROOF. First, note that

$$|\exp(x) - 1| \leq 1 \vee \exp(x),$$

and

$$|\exp(x) - 1| = \left| \int_0^x e^y dy \right| \leq |x|(1 \vee \exp(x)), \quad x \in \mathbf{R}.$$

So we have

$$|\exp(x) - (1 + x)| = \left| \int_0^x (e^y - 1) dy \right| \leq (|x| \wedge |x|^2)(1 \vee \exp(x))$$

for any $x \in \mathbf{R}$. Therefore we see that

$$|\exp(x) - (1 + x)| \leq |x|^2 \exp(|x|), \quad x \in \mathbf{R}.$$

This implies that

$$\begin{aligned} & |E[\exp(sY1_{\{|Y| \leq b\}})] - (1 + E[sY1_{\{|Y| \leq b\}}])| \\ & \leq |s|^2 \exp(|s|b) E[|Y|^2]. \end{aligned}$$

Since

$$|E[sY1_{\{|Y| \leq b\}}]| = |sE[Y, |Y| > b]| \leq |s|b^{-1} E[|Y|^2],$$

we have our assertion. \square

PROPOSITION 24. *Let X be a random variable and assume that*

$$E[|X|^2] < \infty \quad \text{and} \quad E[X] = 0.$$

Then for any $t > 0$ and $n \geq 1$

$$n \log E\left[\exp\left(\pm \frac{1}{tn^{1/2}} X 1_{\{|X| \leq tn^{1/2}\}}\right)\right] \leq \frac{6}{t^2} E[|X|^2].$$

PROOF. Let $Y = (1/t)X$, $s = \pm n^{-1/2}$, $b = n^{1/2}$, and apply Proposition 23. Since $\log(1+x) \leq x$, $x \geq 0$, we have our assertion. \square

Now let X_n , $n = 1, 2, \dots$, be independent identically distributed random variables. Throughout this section we assume that

$$E[|X_1|^2] < \infty \text{ and } E[X_1] = 0.$$

PROPOSITION 25. For any $s, t > 0$ and $\varepsilon > 0$

$$P\left(\left|\sum_{k=1}^n X_k 1_{\{|X_k| \leq tn^{1/2}\}}\right| \geq sn^{1/2}\right) \leq 2 \exp\left(\frac{6}{t^2} E[|X_1|^2]\right) \exp\left(-\frac{s}{t}\right).$$

PROOF. We see that

$$\begin{aligned} & P\left(\pm \sum_{k=1}^n X_k 1_{\{|X_k| \leq tn^{1/2}\}} \geq sn^{1/2}\right) \\ & \leq \exp\left(-\frac{s}{t}\right) E\left[\exp\left(\frac{\pm 1}{tn^{1/2}} \sum_{k=1}^n X_k 1_{\{|X_k| \leq tn^{1/2}\}}\right)\right] \\ & \leq \exp\left(-\frac{s}{t}\right) E\left[\exp\left(\frac{\pm 1}{tn^{1/2}} X_1 1_{\{|X_1| \leq tn^{1/2}\}}\right)\right]^n. \end{aligned}$$

Then by Proposition 24 we have our assertion.

Let $F : \mathbf{R} \rightarrow [0, 1]$ and $\bar{F} : \mathbf{R} \rightarrow [0, 1]$ be given by

$$F(x) = P(X_1 \leq x), \quad x \in \mathbf{R}$$

and

$$\bar{F}(x) = P(X_1 > x), \quad x \in \mathbf{R}.$$

Then we have the following.

PROPOSITION 26. (1) For any $t, s > 0$, and $n \geq 2$,

$$P\left(\left|\sum_{k=2}^n X_k 1_{\{|X_k| \leq tn^{1/2}\}}\right| > sn^{1/2}\right) \leq 2 \exp\left(\frac{6}{t^2} E[|X_1|^2]\right) \exp\left(-\frac{s}{t}\right).$$

(2) For any $s, t > 0$, $\varepsilon \in (0, 1)$ with $t < (1 - \varepsilon)s$,

$$\begin{aligned} & \left|P\left(\sum_{k=1}^n X_k > sn^{1/2}\right) - nP\left(X_1 + \sum_{k=2}^n X_k 1_{\{|X_k| \leq tn^{1/2}\}} > sn^{1/2}\right), \right. \\ & \quad \left. \left|\sum_{k=2}^n X_k 1_{\{|X_k| \leq tn^{1/2}\}}\right| \leq \varepsilon sn^{1/2}\right| \\ & \leq 2n(n-1)(F(-tn^{1/2}) + \bar{F}(tn^{1/2}))^2 + 2 \exp\left(\frac{6}{t^2} E[|X_1|^2]\right) \exp\left(-\frac{s}{t}\right) \\ & \quad + 2n(F(-tn^{1/2}) + \bar{F}(tn^{1/2})) \exp\left(\frac{6}{t^2} E[|X_1|^2]\right) \exp\left(-\frac{\varepsilon s}{2t}\right) \end{aligned}$$

PROOF. Note that

$$\begin{aligned} & P\left(\left|\sum_{k=2}^n X_k 1_{\{|X_k| \leq tn^{1/2}\}}\right| > sn^{1/2}\right) \\ & = P\left(\left|\sum_{k=1}^{n-1} X_k 1_{\{|X_k| \leq \tilde{t}(n-1)^{1/2}\}}\right| > \tilde{s}(n-1)^{1/2}\right), \end{aligned}$$

where

$$\tilde{t} = t\left(\frac{n}{n-1}\right)^{1/2}, \quad \tilde{s} = s\left(\frac{n}{n-1}\right)^{1/2}.$$

So we have the assertion (1) from Proposition 25.

Let us denote

$$\tilde{F}(x) = P(|X_1| > x) \leq F(-x) + \bar{F}(x), \quad x > 0.$$

Note that

$$P\left(\sum_{k=1}^n X_k > sn^{1/2}\right) = \sum_{m=0}^n I_m,$$

where

$$I_m = P\left(\sum_{k=1}^n X_k > sn^{1/2}, \sum_{k=1}^n 1_{\{|X_k| > tn^{1/2}\}} = m\right), \quad m = 0, 1, \dots, n.$$

Then we have

$$I_m = \binom{n}{m} P\left(\sum_{k=1}^n X_k > sn^{1/2}, |X_i| > tn^{1/2}, i = 1, \dots, m, \right. \\ \left. |X_j| \leq tn^{1/2}, j = m+1, \dots, n\right),$$

for $m = 0, 1, \dots, n$. So we see that

$$\sum_{m=2}^n I_m \leq \sum_{m=2}^n \frac{n(n-1)}{m(m-1)} \binom{n-2}{m-2} \tilde{F}(tn^{1/2})^m (1 - \tilde{F}(tn^{1/2}))^{n-m} \\ (15) \quad \leq \frac{n(n-1)}{2} \tilde{F}(tn^{1/2})^2.$$

Also, by Proposition 25, we have

$$(16) \quad I_0 \leq 2 \exp\left(-\frac{s}{t}\right) \exp\left(\frac{6}{t^2} E[|X_1|^2]\right).$$

Let

$$A_1 = \{|X_1| > tn^{1/2}\}, \quad A_2 = \{|X_k| \leq tn^{1/2}, k = 2, 3, \dots, n\},$$

$$B_1 = \left\{X_1 + \sum_{k=2}^n X_k 1_{\{|X_k| \leq tn^{1/2}\}} > sn^{1/2}\right\},$$

and

$$B_2 = \left\{\left|\sum_{k=2}^n X_k 1_{\{|X_k| \leq tn^{1/2}\}}\right| \leq \varepsilon sn^{1/2}\right\}.$$

Note that $B_1 \cap B_2 \subset A_1$, since $t < (1 - \varepsilon)s$. So we see that

$$|P(B_1 \cap A_1 \cap A_2) - P(B_1 \cap B_2)| \\ \leq P(B_1 \cap B_2^c \cap A_1 \cap A_2) + P(B_1 \cap B_2 \cap A_1 \cap A_2^c) \\ (17) \quad \leq P(A_1)P(B_2^c) + P(A_1)P(A_2^c).$$

Note that

$$P(A_2^c) \leq \sum_{k=2}^n P(|X_k| > tn^{1/2}) = (n-1)\tilde{F}(tn^{1/2}).$$

Also, by the assertion (1) we have

$$P(B_2^c) \leq 2 \exp\left(\frac{6}{t^2} E[|X_1|^2]\right) \exp\left(-\frac{\varepsilon s}{2t}\right).$$

Since $I_1 = nP(B_1 \cap A_1 \cap A_2)$, we have the assertion from Equations (15), (16) and (17).

This completes the proof. \square

8. Some Estimates

In this section, we assume that (A-1) and (A-5).

Let $g : (x_0, \infty) \rightarrow \mathbf{R}$, $H : [-1/2, 1/2] \times (2x_0, \infty) \rightarrow (0, \infty)$ and $R : [-1/2, 1/2] \times (2x_0, \infty) \rightarrow (0, \infty)$ be given by

$$g(x) = x^2 \frac{d^2}{dx^2} (\log \bar{F})(x) - \alpha, \quad x > x_0,$$

$$H(y; x) = \frac{\bar{F}(x(1+y))}{\bar{F}(x)}, \quad y \in [-1/2, 1/2], \quad x > 2x_0,$$

and

$$R(y; x) = H(y; x) - \left\{1 - \alpha y + \frac{\alpha(\alpha+1)y^2}{2}\right\}, \quad y \in [-1/2, 1/2], \quad x > 2x_0,$$

We prove the following in this section.

PROPOSITION 27. *There are functions $a : (2x_0, \infty) \rightarrow \mathbf{R}$, $c : (2x_0, \infty) \rightarrow [0, \infty)$ and a constant $C > 0$ such that $a(x) \rightarrow 0$ and $c(x) \rightarrow 0$, as $x \rightarrow \infty$, and that*

$$|R(y; x) - a(x)y| \leq C(c(x)y^2 + |y|^3), \quad y \in [-1/2, 1/2], \quad x > 2x_0.$$

First we prove the following.

PROPOSITION 28. (1) For any $x > x_0$,

$$\frac{d}{dx} \log(x^\alpha \bar{F}(x)) = - \int_x^\infty \frac{g(z)}{z^2}.$$

(2) For any $y \in [-1/2, 1/2]$ and $x > 2x_0$,

$$H(y; x) = (1 + y)^{-\alpha} \exp\left(- \int_0^y dy' \int_{1+y'}^\infty \frac{g(xz)}{z^2} dz\right).$$

PROOF. Note that

$$g(x) = x^2 \frac{d^2}{dx^2} (\log(x^\alpha \bar{F}(x)))$$

and $g(x) \rightarrow 0$ as $x \rightarrow \infty$. Then we see that

$$(18) \quad \frac{d}{dy} (\log(y^\alpha \bar{F}(y))) - \frac{d}{dx} (\log(x^\alpha \bar{F}(x))) = \int_x^y \frac{g(z)}{z^2} dz,$$

and so we see that

$$c_0 = \lim_{y \rightarrow \infty} \frac{d}{dy} (\log(y^\alpha \bar{F}(y)))$$

exists. Note that

$$\exp\left(\int_x^{2x} \frac{d}{dy} (\log(y^\alpha \bar{F}(y))) dy\right) = \frac{L(2x)}{L(x)} \rightarrow 1, \quad x \rightarrow \infty.$$

So we see that $c_0 = 0$. Therefore letting $y \rightarrow \infty$ in Equation (18) we have the assertion (1).

By the assertion (1), we have

$$\frac{d}{dy} \log((1 + y)^\alpha H(y; x)) = -x \int_{x(1+y)}^\infty \frac{g(z)}{z^2} dz = - \int_{1+y}^\infty \frac{g(xz)}{z^2} dz.$$

Since $H(0; x) = 1$, we have the assertion (2). \square

PROPOSITION 29. Let $\tilde{a} : (2x_0, \infty) \rightarrow \mathbf{R}$ and $\tilde{c} : (2x_0, \infty) \rightarrow \mathbf{R}$ be given by

$$\tilde{a}(x) = \frac{d}{dy} ((1 + y)^\alpha H(y, x))|_{y=0},$$

and

$$\tilde{c}(x) = \sup_{y \in [-1/2, 1/2]} \left| \frac{d^2}{dy^2} ((1+y)^\alpha H(y, x)) \right|.$$

Then $\tilde{a}(x) \rightarrow 0$ and $\tilde{c}(x) \rightarrow 0$, as $x \rightarrow \infty$, and that

$$|H(y; x) - (1+y)^{-\alpha} - \tilde{a}(x)y(1+y)^{-\alpha}| \leq 2^\alpha \tilde{c}(x)y^2, \\ y \in [-1/2, 1/2], \quad x > 2x_0.$$

PROOF. By Proposition 11 We have

$$\frac{d}{dy} ((1+y)^\alpha H(y; x)) = -(1+y)^\alpha H(y; x) \int_{1+y}^{\infty} \frac{g(xz)}{z^2} dz$$

and so

$$\tilde{a}(x) = - \int_1^{\infty} \frac{g(xz)}{z^2} dz$$

Similarly, we have

$$\frac{d^2}{dy^2} ((1+y)^\alpha H(y; x)) \\ = (1+y)^\alpha H(y; x) \left\{ \left(\int_{1+y}^{\infty} \frac{g(xz)}{z^2} dz \right)^2 - (1+y)^{-2} g(x(1+y)) \right\}$$

Therefore we have

$$\tilde{c}(x) \leq 2^\alpha \left\{ \left(\int_{1/2}^{\infty} \frac{|g(xz)|}{z^2} dz \right)^2 + 4 \sup\{|g(x(1+y))|\}; \quad y \in [-1/2, 1/2] \right\} \\ \times \exp\left(\int_{1/2}^{\infty} \frac{|g(xz)|}{z^2} dz \right)$$

These imply that $\tilde{a}(x) \rightarrow 0$, $\tilde{c}(x) \rightarrow 0$, as $x \rightarrow \infty$. Also we have

$$|(1+y)^\alpha H(y; x) - (1+\tilde{a}(x))| \leq \tilde{c}(x)y^2, \quad x \geq 2x_0, \quad y \in [-1/2, 1/2].$$

This implies our assestion. \square

Now Proposition 27 is an easy corollary to Proposition 29.

9. Proof of Thoerem 3.

In this section, we assume that X_n , $n = 1, 2, \dots$, are i.i.d. random variables, $\alpha > 2$ and (A-1) - (A-5) are satisfied. Let $p = (\alpha + 2)/2$ and $\beta = (\alpha + p)/2$. Then we see that $E[|X_1|^p] < \infty$ and there is a $C_0 > 1$ such that

$$F(-x) + \bar{F}(x) \leq C_0 x^{-\beta}, \quad x \geq 1.$$

PROPOSITION 30. *Let $b(x) = E[X_1, |X_1| \leq x] = -E[X_1, |X_1| > x]$, $x > 0$. Then we have the following.*

(1) $|b(x)| \leq E[|X_1|^p]^{1/p} (F(-x) + \bar{F}(x))^{1-1/p} \leq C_0 x^{-\beta(p-1)/p} E[|X_1|^p]^{1/p}$, $x \geq 1$.

(2) *There is a constant $C_1 > 1$ only dependent on p such that*

$$\begin{aligned} E\left[\left|\sum_{k=1}^n X_k 1_{\{|X_k| \leq x\}}\right|^p\right]^{1/p} &\leq C_1 n^{1/2} (E[|X_1|^p]^{1/p} + |b(x)|) + n|b(x)| \\ &\leq C_1 E[|X_1|^p]^{1/p} (1 + C_0) (n^{1/2} + nx^{-\beta(p-1)/p}) \end{aligned}$$

for any $n = 1, 2, \dots$, and $x \geq 1$.

PROOF. The assertion (1) is an easy consequence of Hölder's inequality. So we prove the assertion (2). Since $E[X_k 1_{\{|X_k| \leq x\}} - b(x)] = 0$, $k = 1, 2, \dots$, we see by Burkholder-Davis-Gundy's theorem that there is a constant $C_1 > 0$ depending on p only such that

$$E\left[\left|\sum_{k=1}^n (X_k 1_{\{|X_k| \leq x\}} - b(x))\right|^p\right]^{1/p} \leq C_1 E\left[\left|\sum_{k=1}^n (X_k 1_{\{|X_k| \leq x\}} - b(x))^2\right|^{p/2}\right]^{1/p}$$

Then by Hölder's inequality, we have

$$\begin{aligned} E\left[\left|\sum_{k=1}^n (X_k 1_{\{|X_k| \leq x\}} - b(x))\right|^p\right]^{1/p} &\leq C_1 E\left[n^{p/2-1} \sum_{k=1}^n |X_k 1_{\{|X_k| \leq x\}} - b(x)|^p\right]^{1/p} \\ &= C_1 n^{1/2} E\left[|X_1 1_{\{|X_1| \leq x\}} - b(x)|^p\right]^{1/p} \leq C_1 n^{1/2} (E[|X_1 1_{\{|X_1| \leq x\}}|^p]^{1/p} + |b(x)|) \end{aligned}$$

This implies our assertion. \square

Let $a : (2x_0, \infty) \rightarrow \mathbf{R}$ and $c : (2x_0, \infty) \rightarrow [0, \infty)$ be as in Proposition 27. Also, let

$$Y_n(t) = \sum_{k=1}^n X_k 1_{\{|X_k| \leq tn^{1/2}\}}, \quad n \geq 2, t > 0.$$

Then we have the following.

PROPOSITION 31. *Let $r \in ((\alpha + 2)/(2\alpha), 1)$. Then for any $\varepsilon \in (0, 1/2)$*

$$\begin{aligned} \overline{\lim}_{n \rightarrow \infty} \sup \{ & s^2 E[|H(-\frac{1}{sn^{1/2}} Y_n(t), sn^{1/2}), \\ & |Y_n(t)| \leq \varepsilon sn^{1/2}] - (1 + \frac{\alpha(\alpha + 1)}{2s^2})]; \\ & s \geq (\log n)^{1/2}, t \geq (\log n)^{-1} s^{(1+r)/2} \} = 0. \end{aligned}$$

PROOF. Let $s \geq (\log n)^{1/2}$, $t \geq (\log n)^{-1} s^{(1+r)/2}$, and $n \geq 3$. Then $tn^{1/2} \geq 1$. Note that

$$r\beta(p-1)/p > 1 + \frac{3(\alpha-2)}{8} > 1.$$

We see that

$$\begin{aligned} & E[H(-\frac{1}{sn^{1/2}} Y_n(t), sn^{1/2}), |Y_n(t)| \leq \varepsilon sn^{1/2}] - (1 + \frac{\alpha(\alpha + 1)}{2s^2}) \\ &= \frac{a(sn^{1/2}) - \alpha}{sn^{1/2}} E[Y_n(t)] + \frac{\alpha(\alpha + 1)}{2s^2 n} (E[Y_n(t)^2] - n) \\ &- E[1 + \frac{a(sn^{1/2}) - \alpha}{sn^{1/2}} Y_n(t) + \frac{\alpha(\alpha + 1)}{2s^2 n} Y_n(t)^2, |(sn^{1/2})^{-1} Y_n(t)| > \varepsilon] \\ &+ E[R(-\frac{1}{sn^{1/2}} Y_n(t), sn^{1/2}), |(sn^{1/2})^{-1} Y_n(t)| \leq \varepsilon]. \end{aligned}$$

Note that

$$\begin{aligned} s|E[Y_n(t)]| &= ns|b(tn^{1/2})| \leq C_0 E[|X_1|^p]^{1/p} s (tn^{1/2})^{-\beta(p-1)/p} \\ &\leq C_0 E[|X_1|^p]^{1/p} (n^{1/2} (\log n)^{-1})^{-\beta(p-1)/p} s^{1-(1+r)\beta(p-1)/2p}, \\ E[Y_k(t)^2] - n &= n(E[(X_1 1_{|X_1| \leq tn^{1/2}})^2] - b(tn^{1/2})^2) + E[Y_n(t)]^2 - n \end{aligned}$$

$$= -nE[X_1^2, |X_1| > tn^{1/2}] + n(n-1)b(tn^{1/2})^2,$$

and

$$n^{-p/2}E[|Y_n(t)|^p] \leq C_1^p(1+C_0)^pE[|X_1|^p](1+t^{-\beta(p-1)/p})n^{1/2(1-\beta(p-1)/p)p}.$$

So we see that

$$\begin{aligned} \frac{1}{n}|E[Y_k(t)^2] - n| &\leq E[X_1^2, |X_1| > tn^{1/2}] \\ &+ C_0(n-1)n^{-\beta(p-1)/p}(\log n)^{2\beta(p-1)/p}E[|X_1|^p]^{2/p}, \\ &s^2(sn^{1/2})^{-k}E[|Y_n(t)|^k, |(sn^{1/2})^{-1}Y_n(t)| > \varepsilon] \\ &\leq s^{2-p}\varepsilon^{-p+k}n^{-p/2}E[|Y_n(t)|^p], \quad k = 0, 1, 2, \end{aligned}$$

and

$$\begin{aligned} s^2E[|R(-\frac{1}{sn^{1/2}}Y_n(t), sn^{1/2}), |(sn^{1/2})^{-1}Y_n(t)| \leq \varepsilon], \\ \leq C_2(|c(sn^{1/2})|n^{-1}E[Y_n(t)^2] + \varepsilon^{-p}s^{2-p}n^{-p/2}E[|Y_n(t)|^p]). \end{aligned}$$

Combining them, we have our assertion. \square

Now we prove Theorem 3. Let $\beta : \mathbf{N} \rightarrow (0, \infty)$ be such that

$$\frac{\beta(n)}{(\log n)^{1/2}} \rightarrow \infty, \quad n \rightarrow \infty.$$

Assume that Theorem 3 is not valid. Then there is a sequence of positive numbers $\{s'_n\}_{n=1}^\infty$ such that $s'_n \geq n^{1/2}\beta(n)$, $n = 1, 2, \dots$, and

$$\overline{\lim}_{n \rightarrow \infty} \frac{(s'_n)^2}{n} \left| \frac{P(\sum_{k=1}^n X_k > s'_n)}{n\bar{F}(s'_n)} - \left(1 + \frac{\alpha(\alpha+1)vn}{(s'_n)^2}\right) \right| > 0.$$

Let $s_n = n^{-1/2}s'_n \geq \beta(n)$. Let us take an $r \in ((\alpha+2)/(2\alpha), 1)$ and fix it. Let t_n , $n = 1, 2, \dots$, be a sequence of positive numbers given by

$$t_n = (\log n)^{-1/2} + (\log n)^{-1}s_n^{(1+r)/2}, \quad n \geq 2.$$

Then we have the following.

$$(19) \quad t_n s_n \geq \frac{\beta(n)}{(\log n)^{1/2}} \rightarrow \infty, \quad n \rightarrow \infty,$$

$$(20) \quad \frac{2s_n}{t_n} \geq ((\log n)^{1/2} s_n) \wedge ((\log n) s_n^{(1-r)/2}), \quad n \geq 2,$$

$$(21) \quad \frac{2s_n}{(\log n)t_n} \geq \left(\frac{\beta(n)}{(\log n)^{1/2}} \right) \wedge (s_n^{(1-r)/2}) \rightarrow \infty, \quad n \rightarrow \infty,$$

and

$$(22) \quad \frac{(t_n n^{1/2})^2}{(s'_n)^{1+r}} \geq \frac{(\log n)^{-2} s_n^{1+r} n}{s_n^{1+r} n^{(1+r)/2}} \rightarrow \infty, \quad n \rightarrow \infty.$$

Therefore by Equation (22), we have

$$\frac{(F(-t_n n^{1/2}) + \bar{F}(t_n n^{1/2}))^2}{\bar{F}(s'_n)^{2r}} \rightarrow 0, \quad n \rightarrow \infty.$$

Since $2r - 1 > 2/\alpha$, we have

$$(23) \quad (s'_n)^2 \frac{(F(-t_n n^{1/2}) + \bar{F}(t_n n^{1/2}))^2}{\bar{F}(s'_n)} \rightarrow 0, \quad n \rightarrow \infty.$$

Also, by Equations (19), (20) and (21) we see that for any $m \geq 1$

$$\begin{aligned} & (ns'_n)^m \exp\left(\frac{m}{t_n^2} - \frac{1}{m} \frac{s_n}{t_n}\right) \\ &= \exp\left(\frac{m}{t_n^2} \left(1 - \frac{1}{3m^2} t_n s_n\right)\right) n^{3m/2} \exp(-(\log n) \frac{1}{3m} \frac{s_n}{(\log n)t_n}) \\ & \quad \times (s_n)^m \exp\left(-\frac{1}{3m} \frac{s_n}{t_n}\right) \rightarrow 0, \quad n \rightarrow \infty. \end{aligned}$$

Note that

$$\begin{aligned} & \frac{(n+1)P(X_{n+1} + Y_n(t_{n+1}) > s_{n+1}(n+1)^{1/2}, |Y_n(t_{n+1})| \leq \varepsilon s_{n+1}(n+1)^{1/2})}{(n+1)\bar{F}(s_{n+1}(n+1)^{1/2})} \\ &= E\left[H\left(-\frac{1}{s_{n+1}(n+1)^{1/2}} Y_n(t_{n+1}), s_{n+1}(n+1)^{1/2}, |Y_n(t_{n+1})| \leq \varepsilon s_{n+1}(n+1)^{1/2}\right)\right]. \end{aligned}$$

From Proposition 31, we see that

$$\begin{aligned} & s_{n+1}^2 |E\left[H\left(-\frac{1}{s_{n+1}(n+1)^{1/2}} Y_n(t_{n+1}), s_{n+1}(n+1)^{1/2}, \right.\right. \\ & \quad \left.\left. |Y_n(t_{n+1})| \leq \varepsilon s_{n+1}(n+1)^{1/2}\right)\right] \\ & \quad \left. - \left(1 + \frac{\alpha(\alpha+1)(n+1)}{s_{n+1}^2(n+1)}\right)\right| \rightarrow 0 \end{aligned}$$

as $n \rightarrow \infty$, by letting $s = s_{n+1}(1 + 1/n)^{1/2}$, $t = t_{n+1}$. Then from this, Proposition 26(2), Equations (19), (20), (21), (22) and (23), we have

$$\frac{(s'_n)^2}{n} \left| \frac{P(\sum_{k=1}^n X_k > s'_n)}{n\bar{F}(s'_n)} - \left(1 + \frac{\alpha(\alpha + 1)n}{(s'_n)^2}\right) \right| \rightarrow 0$$

as $n \rightarrow \infty$. This is a contradiction.

This proves Theorem 3.

10. Proof of Theorem 4

Let $\hat{F}_n : [1, \infty) \rightarrow [0, 1]$, $n \geq 1$, be given by

$$\hat{F}_n(s) = \int_{-\infty}^s \bar{F}((s-x)n^{1/2})\Phi_1(x)dx.$$

Then we have the following.

PROPOSITION 32. *Let $\beta : \mathbf{N} \rightarrow (0, \infty)$ be such that*

$$\frac{\beta(n)}{(\log n)^{1/2}} \rightarrow \infty, \quad n \rightarrow \infty.$$

Then

$$\sup_{s \geq \beta(n)} s^2 \left| \frac{\hat{F}_n(s)}{\bar{F}(n^{1/2}s)} - \left(1 + \frac{\alpha(\alpha + 1)n}{s^2}\right) \right| \rightarrow 0, \quad n \rightarrow \infty.$$

PROOF. By Proposition 27, we see that

$$\begin{aligned} & \left| \frac{\hat{F}_n(s)}{\bar{F}(sn^{1/2})} - \left(1 + \frac{\alpha(\alpha + 1)}{s^2}\right) \right| \\ & \leq \int_{[-s/2, s/2]} |R(y/s; sn^{1/2}) - a(sn^{1/2})(y/s)|\Phi_1(y)dy \\ & + \int_{[-s/2, s/2]^c} 4(1 + (|a(sn^{1/2})| + \alpha(\alpha + 1))\left(\frac{y^2}{s^2}\right)\Phi_1(y)dy. \end{aligned}$$

This and Proposition 27 imply our assertion. \square

It is well known (e.g. Williams [8]) that there is a $C_0 > 0$ such that

$$(24) \quad |\Phi_k(x)| \leq C_0(1+x)^{k-1}\Phi_1(x), \quad x \geq 0, \quad k = 1, \dots, 3K,$$

and

$$(25) \quad C_0^{-1}\Phi_1(x) \leq x\hat{\Phi}_0(x) \leq C_0\Phi_1(x), \quad x \geq 1.$$

Let

$$H_0(n, s) = \Phi_0(s) + n\hat{F}_n(s),$$

$$A_1(n, s) = \sum_{k=1}^2 \frac{n^{-(k-2)/2}}{k!} \Phi_k(s) \int_0^\infty x^k \mu(dx),$$

and

$$A(n, s) = n\hat{F}_n(s) - A_1(n, s).$$

First we prove the following.

PROPOSITION 33.

$$\sup_{s \in [1, \log n]} \frac{|A(n, s) - n\bar{F}(n^{1/2}s)|}{H_0(n, s)} \rightarrow 0, \quad n \rightarrow \infty.$$

PROOF. Let us take a $\gamma \in (0, (\alpha - 2)/(4\alpha))$ and fix it. Let $s \geq 0$ and $n \geq 3$. Note that

$$\hat{F}_n(s) = \sum_{k=1}^4 I_k(n, s),$$

where

$$I_1(n, s) = \int_{s-n^{-\gamma}}^s \bar{F}((s-x)n^{1/2})\Phi_1(x)dx,$$

$$I_2(n, s) = \int_{-s}^{7s/8} \bar{F}((s-x)n^{1/2})\Phi_1(x)dx,$$

$$I_3(n, s) = \int_{7s/8}^{s-n^{-\gamma}} \bar{F}((s-x)n^{1/2})\Phi_1(x)dx,$$

and

$$I_4(n, s) = \int_{-\infty}^{-s} \bar{F}((s-x)n^{1/2})\Phi_1(x)dx.$$

It is easy to see that

$$I_1(n, s) = n^{-1/2} \int_0^{n^{1/2-\gamma}} \bar{F}(y) \Phi_1(s - n^{-1/2}y) dy.$$

Let

$$R(n, s, y) = \Phi_1(s - n^{-1/2}y) - (\Phi_1(s) + n^{-1/2}y\Phi_2(s))$$

Then for $y \in [0, sn^{1/2-\gamma}]$

$$\begin{aligned} |R(n, s, y)| &\leq n^{-1}y^2 \sup_{z \in [s-n^{-\gamma}, s]} |\Phi_3(z)| \\ &\leq C_0 n^{-1}y^2(1+s)^2 \Phi_1(s - n^{-\gamma}) = C_0 n^{-1}y^2(1+s)^2 \Phi_1(s) \exp(sn^{-\gamma} - n^{-2\gamma}/2) \\ &\leq C_0^2 n^{-1}y^2(1+s)^3 \exp(n^{-\gamma}s) \Phi_0(s). \end{aligned}$$

So we see that

$$\begin{aligned} n|I_1(n, s) - \sum_{k=1}^2 \frac{n^{-k/2}}{k!} \Phi_k(s) \int_0^\infty x^k \mu(dx)| \\ \leq C_0^2(1+s)^3 n^{-1/2} \exp(n^{-\gamma}s) \left(\int_0^{n^{1/2-\gamma}} y^2 \bar{F}(y) dy \right) \Phi_0(s) \\ + C_0^2(1+s)n^{1/2} \left(\int_{n^{1/2-\gamma}}^\infty \bar{F}(y) dy \right) \Phi_0(s) + C_0^2(1+s)^2 \left(\int_{n^{1/2-\gamma}}^\infty y \bar{F}(y) dy \right) \Phi_0(s) \end{aligned}$$

This implies that

$$(26) \quad \sup_{s \in [1, \log n]} \Phi_0(s)^{-1} |nI_1(n, s) - \sum_{k=1}^2 \frac{n^{-(k-2)/2}}{k!} \Phi_k(s) \int_0^\infty x^k \mu(dx)| \rightarrow 0, \\ n \rightarrow \infty.$$

Note that

$$I_2(n, s) = \bar{F}(sn^{1/2}) \int_{-s}^{7s/8} \left(1 - \frac{x}{s}\right)^{-\alpha} \frac{L((s-x)n^{1/2})}{L(sn^{1/2})} \Phi_1(x) dx$$

It is easy to see that

$$\sup_{s \in [(\log n)^{1/4}, \log n]} \left| \int_{-s}^{7s/8} \left(1 - \frac{x}{s}\right)^{-\alpha} \frac{L((s-x)n^{1/2})}{L(sn^{1/2})} \Phi_1(x) dx - 1 \right| \rightarrow 0, \quad n \rightarrow \infty$$

Also we see that

$$n|I_2(n, s)| \leq n\bar{F}(sn^{1/2})8^\alpha \int_{-s}^{7s/8} \frac{L((s-x)n^{1/2})}{L(sn^{1/2})} \Phi_1(x) dx$$

Therefore we have

$$\sup_{s \in [1, (\log n)^{1/4}]} \Phi_0(s)^{-1} (n|I_2(n, s)| + n\bar{F}(sn^{1/2})) \rightarrow 0, \quad n \rightarrow \infty.$$

Thus we have

$$(27) \quad \sup_{s \in [1, \log n]} H_0(n, s)^{-1} |nI_2(n, s) - n\bar{F}(sn^{1/2})| \rightarrow 0, \quad n \rightarrow \infty.$$

Note that $\sqrt{3}/2 \leq 7/8$. Then we have

$$\Phi_1(7s/8) \leq (\Phi_1(s))^{3/4},$$

and so we have

$$\begin{aligned} nI_3(n, s) &\leq ns\bar{F}(n^{1/2-\gamma})\Phi_1(7s/8) \\ &\leq (n\bar{F}(n^{1/2} \log n))^{1/2} (s\Phi_1(s))^{3/4} \frac{ns^{1/4}\bar{F}(n^{1/2-\gamma})}{(n\bar{F}(n^{1/2} \log n))^{1/2}}. \end{aligned}$$

Since

$$\sup_{n \geq 3} \sup_{s \in [1, \log n]} \frac{ns^{1/4}\bar{F}(n^{1/2-\gamma})}{(n\bar{F}(n^{1/2} \log n))^{1/2}} < \infty,$$

we see that there is a constant $C > 0$ such that

$$\begin{aligned} nI_3(n, s) &\leq C(n\bar{F}(sn^{1/2}))^{1/2} \Phi_0(s)^{3/4} \leq C(n\bar{F}(sn^{1/2}))^{1/4} H_0(n, s), \\ &n \geq 3, \quad s \in [1, \log n]. \end{aligned}$$

So we have

$$(28) \quad \sup_{s \in [1, \log n]} H_0(n, s)^{-1} |nI_3(n, s)| \rightarrow 0, \quad n \rightarrow \infty.$$

Also we have

$$n|I_4(n, s)| \leq n\bar{F}(2sn^{1/2})\Phi_0(s).$$

So this equation, Equations (26) (27) and (28) imply our assertion. \square

PROPOSITION 34. (1) *There is a $C > 0$ such that*

$$\Phi_0(s) + |A_1(n, s)| \leq Cn^{-2}\bar{F}(n^{1/2}s), \quad n \geq 2, \quad s \geq \log n.$$

(2)

$$\sup_{s \in [1, \infty)} \frac{|A(n, s) - n\bar{F}(n^{1/2}s)|}{H_0(n, s)} \rightarrow 0, \quad n \rightarrow \infty.$$

PROOF. The assertion (1) is obvious from Equations (24) and (25). To prove the assertion (2), because of Proposition 33, it is sufficient to prove

$$\sup_{s \in [\log n, \infty)} \frac{|A(n, s) - n\bar{F}(n^{1/2}s)|}{H_0(n, s)} \rightarrow 0, \quad n \rightarrow \infty.$$

However, by Theorem 3 and Proposition 32, we see that

$$\sup_{s \in [\log n, \infty)} \left| \frac{n\hat{F}(n^{1/2}s)}{n\bar{F}(n^{1/2}s)} - 1 \right| \rightarrow 0, \quad n \rightarrow \infty.$$

Therefore, combining this with the assertion (1), we have the assertion (2). \square

Now let us prove Theorem 4.

By Proposition 34(2), we see that

$$\sup_{s \in [1, \infty)} \left| \frac{H(n, s)}{H_0(n, s)} - 1 \right| \rightarrow 0, \quad n \rightarrow \infty.$$

Therefore there is an n_0 such that

$$H(n, s) \geq \frac{1}{2}H_0(n, s), \quad n \geq n_0, \quad s \geq 1.$$

By Equation (24), we see that there is a $C > 0$ such that

$$\sup_{s \in [1, n^{1/12}]} |G(n, s) - H(n, s)| \leq Cn^{-1/12}, \quad n \geq 1$$

Then combining this with Theorem 2, we see that there are $C > 0$ and $\delta_0 \in (0, 1/12)$ such that

$$\sup_{s \in [1, n^{\delta_0}]} \left| \frac{P(\sum_{k=1}^n X_k > n^{1/2}s)}{H(n, s)} - 1 \right| \leq Cn^{-\delta_0}, \quad n \geq n_0.$$

On the other hand, by Theorem 3 and Proposition 32 we see that there is a $C > 0$ such that

$$\sup_{s \in [n^{\delta_0}, \infty)} \left| \frac{P(\sum_{k=1}^n X_k > n^{1/2}s)}{n\hat{F}_n(s)} - 1 \right| \leq Cn^{-2\delta_0}, \quad n \geq n_0.$$

So we see by Proposition 34 that we see that there is a $C > 0$ such that

$$\sup_{s \in [n^{\delta_0}, \infty)} \left| \frac{P(\sum_{k=1}^n X_k > n^{1/2}s)}{H(n, s)} - 1 \right| \leq Cn^{-2\delta_0}, \quad n \geq n_0.$$

These imply Theorem 4.

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